## An improved estimate on the degree of approximation to functions of bounded variation by certain operators

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ABSTRACT. In [5] Guo introduced a modification  $\hat{M}_n(f,x)$  of the Meyer-König-Zeller operators  $M_n(f,x)$  and gave an estimate on the rate of convergence to f of the modified operators when f is a function of bounded variation. In the present paper we establish a sharp estimate on the operators  $\hat{M}_n(f,x)$  using some results in probability theory.

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#### ditweed as we make the 1. Introduction

For the Meyer-König-Zeller operators (see [3]) one has

$$M_n(f,x) = \sum_{k=0}^{\infty} p_{n,k}(x) f\left(\frac{k}{n+k}\right), \quad p_{n,k}(x) = \binom{n+k-1}{k} x^k (1-x)^n.$$

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Recently, Guo [5] introduced an integral modification of these operators by

$$\hat{M}_n(f,x) = \sum_{k=1}^{\infty} p_{n,k+1}(x) \frac{(n+k-2)(n+k-3)}{(n-2)} \int_0^1 p_{n-2,k-1}(t) f(t) dt. \quad (1.1)$$

Cheng [2] studied the rate of convergence of  $M_n(f,x)$  to f(x) when f is a function of bounded variation. Guo ([4], [5]) estimated the rate of convergence of Durrmeyer operators and integrated Meyer-König-Zeller operators  $M_n(f,x)$  using a probabilistic approach.

Guo [5] has also given an estimate for the rate of convergence to f of  $M_n(f, x)$  when f is function of bounded variation, as below

Theorem 1.1. Let f be a function of bounded variation of [0, 1],  $x \in (0, 1)$ . Then, for n sufficiently large,

$$|\hat{M}_n(f,x) - \frac{1}{2} \{ f(x+) + f(x-) \} | \le \frac{7}{nx} \sum_{k=1}^n V_{x-x/\sqrt{k}}^{x+(1-x)/\sqrt{k}} (g_x) + \frac{50}{\sqrt{n} x^{3/2}} |f(x+) - f(x-)|,$$

where

$$g_x(t) = \left\{ egin{array}{ll} f(t) - f(x+), & x < t \leq 1 \ 0, & t = x \ f(t) - f(x-), & 0 \leq t < x, \end{array} 
ight.$$

and  $V_a^b(g_x)$  is the total variation of  $g_x$  on [a,b].

In the present paper we establish a sharper estimate than the above by using some results in probability theory.

### 2. Auxiliary results

The following results wil be needed to prove our main theorem.

Lemma 2.1. [1, p. 104 & 110]. (Berry-Esseen Theorem). Let  $X_1, X_2, \dots, X_n$  be n independent and identically distributed (i. i. d) random variables with zero mean and a finite absolute third moment. If  $\rho_2 \equiv E(X_1^2) > 0$ , then

$$\sup_{x \in R} |F_n(x) - \varphi(x)| \le (0.409)\ell_{3,n},$$

where  $F_n$  is the distribution function of  $(X_1 + X_2 + \cdots + X_n)n\rho_2$ ,  $\varphi$  is the standard normal distribution function and  $\ell_{3,n}$  is the Liapounov ratio given by

$$\ell_{3,n} = \left(\frac{\rho_3}{\rho_2^{3/2}}\right) n^{-1/2} \quad , \quad \rho_3 = E(|X_1|^3).$$

**Lemma 2.2** [6]. If  $\{\xi_i\}(i=1,2,\cdots)$  are independent random variables with the same geometric distribution

$$p(\xi_i = k) = x^k(1-x), x \in (0,1), i = 1,2,...,$$

then

$$E(\xi_i) = \frac{x}{1-x}$$
,  $\rho_2 = E(\xi_i - E(\xi_i))^2 = \frac{x}{(1-x)^2}$ 

and  $\eta_n = \sum_{i=1}^n \xi_i$  is a random variable with distribution

$$P(\eta_n = k) = \binom{n+k-1}{k} x^k (1-x)^n.$$
 (2.1)

Lemma 2.3. If y is a positive valued random variable with a non-degenerate probability distribution then (by Schwarz's inequality)

$$E(y^3) \le (E(y^2).E(y^4))^{1/2},$$

provided  $E(y^2)$ ,  $E(y^3)$  and  $E(y^4) < \infty$ .

**Lemma 2.4.** For  $k \in N$  and  $x \in (0,1)$ ,

$$p_{n,k}(x) \le 19/(6\sqrt{n} \ x^{3/2}), \quad n = 1, 2, \dots$$
 (2.2)

Proof. By (2.1) we have

$$p_{n,k}(x) = P(\eta_n = k) = P(k - 1 < \eta_n \le k)$$

$$= P\left(\frac{k - 1 - nx/(1 - x)}{\sqrt{nx}/(1 - x)} < \frac{\eta_n - nx/(1 - x)}{\sqrt{nx}/(1 - x)} \le \frac{k - nx/(1 - x)}{\sqrt{nx}/(1 - x)}\right),$$

and using Lemma 2.1 we obtain

$$\left| P(\eta_n = k) - \frac{1}{\sqrt{2\pi}} \int_{\frac{k-1-nx/(1-x)}{\sqrt{nx}/(1-x)}}^{\frac{k-nx/(1-x)}{\sqrt{nx}/(1-x)}} e^{-t^2/2} dt \right| < \frac{2(0.409)\rho_3}{\sqrt{n} \rho_2^{3/2}}.$$
 (2.3)

Next, we estimate  $\rho_3$ . By an easy computation we have

$$\begin{cases}
T_0(x) &= \sum_{k=0}^{\infty} x^k (1-x) = 1, \\
T_1(x) &= \sum_{k=0}^{\infty} kx^k (1-x) = \frac{x}{1-x}, \\
T_2(x) &= \sum_{k=0}^{\infty} k^2 x^k (1-x) = \frac{x(1+x)}{(1-x)^2}, \\
T_3(x) &= \sum_{k=0}^{\infty} k^3 x^k (1-x) = \frac{x^3 + 4x^2 + x}{(1-x)^3}, \\
T_4(x) &= \sum_{k=0}^{\infty} k^4 x^k (1-x) = \frac{x^4 + 11x^3 + 11x^2 + x}{(1-x)^4}.
\end{cases} (2.4)$$

Also, if  $M_r(x) (\equiv \sum_{k=0}^{\infty} (k - \frac{x}{1-x})^r x^k (1-x))$  stands for the central moment of order r about the mean  $\frac{x}{1-x}$ , it is easily checked, by using (2.4), that

$$M_2(x) = \sum_{j=0}^2 \binom{2}{j} T_{2-j}(x) (T_1(x))^j (-l)^j = rac{x}{(1-x)^2}$$

and that

$$M_4(x) = \sum_{j=0}^4 {4 \choose j} T_{4-j}(x) (T_1(x))^j (-l)^j = \frac{x^3 + 7x^2 + x}{(1-x)^4}.$$

Thus, in view of Lemma 2.3, we have

$$\rho_3 \le (M_2(x) \cdot M_4(x))^{1/2} = \left(\frac{x}{(1-x)^2} \cdot \frac{x^3 + 7x^2 + x}{(1-x)^4}\right)^{1/2}$$

$$\le \frac{3}{(1-x)^3}, \quad x \in (0,1).$$

So, the right hand side of (2.3) is less than  $8/(3\sqrt{n} \ x^{3/2})$ , and the absolute value of the second term on the left hand side of (2.3) is less than  $(1-x)/\sqrt{2\pi nx}$ , i.e., than  $1/(2\sqrt{n} \ x^{3/2})$  (as  $1/\sqrt{2\pi} < 1/2$  and  $x \in (0,1)$ ).

Hence

$$p_{n,k}(x) = P(\eta_n = k) \le \frac{19}{6\sqrt{n} \ x^{3/2}}.$$

**Lemma 2.5** [5]. For  $k \ge 1, x \in (0,1)$  and n > 2, we have

$$\sum_{j=0}^{k-1} p_{n-1,j}(x) = \frac{(n+k-2)(n+k-3)}{n-2} \int_x^1 p_{n-2,k-1}(t) dt.$$

Lemma 2.6. For  $k \geq 1, x \in (0,1)$  and n > 2, we have

$$\left| \sum_{j=2}^{k+1} p_{n,j}(x) - \sum_{j=1}^{k-1} p_{n-1,j}(x) \right| \le 19/(6\sqrt{n} \ x^{3/2}).$$

Proof. Since

$$\sum_{j=0}^{k+1} p_{n,j}(x) = P(\eta_n \le k+1) \quad \text{and} \quad \sum_{j=0}^{k-1} p_{n-1,j}(x) = P(\eta_{n-1} \le k-1),$$

Lemma 2.1 gives

$$\left| \sum_{j=0}^{k+1} p_{n,j}(x) - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{k+1-nx/(1-x)}{\sqrt{nx}/(1-x)}} e^{-t^2/2} dt \right| < \frac{4}{3\sqrt{n} \ x^{3/2}}$$

$$\left| \sum_{j=0}^{k-1} p_{n-1,j}(x) - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\frac{k-1-(n-1)x/(1-x)}{\sqrt{(n-1)x/(1-x)}}} e^{-t^2/2} dt \right| < \frac{4}{3\sqrt{(n-1)}x^{3/2}}.$$

On the other hand,

$$p_{n,0}(x) + p_{n,1}(x) = o\left(\frac{1}{n}\right).$$

Hence, for n sufficiently large we have

$$\left| \sum_{j=2}^{k+1} p_{n,j}(x) - \sum_{j=0}^{k-1} p_{n-1,j}(x) \right| < \frac{8}{3\sqrt{n} \ x^{3/2}} + \frac{1}{\sqrt{2\pi}} \int_{I_{n,k}} e^{-t^2/2} dt$$

$$\leq \frac{19}{6\sqrt{n} \ x^{3/2}}$$

where  $I_{n,k} = \left[\frac{k-1-nx/(1-x)}{\sqrt{nx}/(1-x)}, \frac{k-nx/(1-x)}{\sqrt{nx}/(1-x)}\right]$ , and the lemma is proved.

**Lemma 2.7** [5]. For all  $x \in (0,1)$  and n sufficiently large we have

$$\hat{M}_n((t-x)^2,x) = rac{2x(1-x)^2}{n-2} + o\left(rac{1}{n^2}
ight),$$
  $\hat{M}_n((t-x)^4,x) \leq rac{8}{n^2}.$ 

**Lemma 2.8.** Let  $K_n(x,t) = \sum_{k=1}^{\infty} \frac{(n+k-2)(n+k-3)}{n-2} p_{n,k+1}(x) p_{n-2,k-1}(x)$ . If n is sufficiently large, then:

(i) For  $0 \le y < x$ ,

$$\int_0^y K_n(x,t)dt \leq \frac{3x(1-x)^2}{n(x-y)^2}.$$

(ii) For  $x < z \le 1$ ,

$$\int_{y}^{1} K_{n}(x,t)dt \leq \frac{3x(1-x)^{2}}{n(y-x)^{2}}.$$

The proof of this lemma easily follows from Lemma 2.7.

#### 3. Main result for the masking the smart

In this section, we state and prove the following theorem:

Theorem 3.1. Let f be a function of bounded variation on  $[0,1], x \in (0,1)$ . Then, for sufficiently large n, we have

$$\left| \hat{M}_n(f, x) - \frac{1}{2} \{ f(x+) + f(n-) \} \right| \le \frac{7}{nx} \sum_{k=1}^n V_{x-x/\sqrt{k}}^{x+x/\sqrt{k}}(g_x) + \frac{4.75}{\sqrt{n} \ x^{3/2}} |f(x+) - f(x-)|,$$

where  $g_x(t)$  and  $V_a^b(g_x)$  are defined as in Theorem 1.1.

Proof. First

$$\left| \hat{M}_{n}(f,x) - \frac{1}{2} \{ f(x+) + f(x-) \} \right| \\ \leq \hat{M}_{n}(g_{x},x) + \frac{1}{2} |f(x+) - f(x-)| \cdot |\hat{M}_{n}(\operatorname{Sign}(t-x),x)|$$

Thus, to estimate the left hand side we need estimates for  $\hat{M}_n(g_x, x)$  and  $\hat{M}_n(\operatorname{Sign}(t-x), x)$ .

To estimate  $\hat{M}_n(\operatorname{Sign}(t-x), x)$ , we first decompose it into two parts as follows:

$$\hat{M}_n(\operatorname{Sign}(t-x),x) = \int_0^1 \operatorname{Sign}(t-x)K_n(x,t)dt$$

$$= \int_x^1 K_n(x,t)dt - \int_0^x K_n(x,t)dt = A_n(x) - B_n(x), \text{ say.}$$

. Making use of Lemma 2.5, we have

$$A_n(x) = \int_x^1 K_n(x,t)dt$$

$$= \sum_{k=1}^\infty p_{n,k+1}(x) \frac{(n+k-2)(n+k-3)}{n-2} \int_x^1 p_{n-2,k-1}(t)dt$$

$$= \sum_{k=1}^\infty \left( p_{n,k+1}(x) \sum_{j=0}^{k-1} p_{n-1,j}(x) \right).$$

Now, from Lemma 2.6, it follows that

$$\left| A_n(x) - \sum_{k=1}^{\infty} \left( p_{n,k+1}(x) \sum_{j=2}^{k+1} p_{n,j}(x) \right) \right| \le 19/(6\sqrt{n} \ x^{3/2}).$$

Let 
$$S = \sum_{k=2}^{\infty} \left( p_{n,k}(x) \sum_{j=2}^{k} p_{n,j}(x) \right)$$
  

$$= p_{n,2}p_{n,2} + p_{n,3}(p_{n,2} + p_{n,3}) + \dots + p_{n,m}(p_{n,2} + p_{n,3} + \dots + p_{n,m}) + \dots$$

Since

$$(p_{n,2}+p_{n,3}+\cdots+p_{n,m}+\cdots)(p_{n,2}+p_{n,3}+\cdots+p_{n,m}+\cdots)$$

$$=p_{n,2}(p_{n,2}+\cdots+p_{n,m}+\cdots)+\cdots+p_{n,m}(p_{n,2}+\cdots+p_{n,m}+\cdots)+\cdots$$

$$=1+o\left(\frac{1}{n}\right),$$

then

$$p_{n,2}(p_{n,3} + p_{n,4} + \cdots) + p_{n,3}(p_{n,4} + p_{n,5} + \cdots) + \cdots + p_{n,m}(p_{n,m+1} + p_{n,m+2} + \cdots) + \cdots$$

$$= p_{n,3}p_{n,2} + p_{n,4}(p_{n,2} + p_{n,3}) + p_{n,5}(p_{n,2} + p_{n,2} + p_{n,3} + p_{n,4}) + \cdots$$

$$= 1 - S + o\left(\frac{1}{n}\right),$$

Now, we can easily obtain that

$$p_{n,2}^2 + p_{n,3}^2 + \cdots + p_{n,m}^2 + \cdots = 2S - 1 + o\left(\frac{1}{n}\right),$$

and using Lemma 2.4, that

$$\left| S - \frac{1}{2} \right| \le \frac{19}{12\sqrt{n} \ x^{3/2}} \sum_{k=2}^{\infty} p_{n,k}(x) + o\left(\frac{1}{n}\right) \le \frac{19}{12\sqrt{n} \ x^{3/2}}.$$

Hence

$$\left| A_n(x) - \frac{1}{2} \right| \le \frac{19}{4\sqrt{n} \ x^{3/2}}.$$

On the other hand,

$$A_n(x) + B_n(x) = \int_0^1 K_n(x, t) dt = 1 + o\left(\frac{1}{n}\right)$$

Therefore,

$$|A_n(x) - B_n(x)| = \left| 2A_n(x) - 1 + o\left(\frac{1}{n}\right) \right| \le \frac{19}{2\sqrt{n} \ x^{3/2}}.$$
 (3.2)

Since the estimate of  $\hat{M}_n(g_x, x)$  is exactly the same as given in [5], we thus have, by using Lemma 2.8, that

$$|\hat{M}_n(g_x, x)| \le \frac{7}{nx} \sum_{k=1}^n V_{x-x/\sqrt{k}}^{x+(1-x)/\sqrt{k}}(g_x). \tag{3.3}$$

Combining the estimates (3.1)-(3.3), the theorem follows.

Remark. It may be shown paralleling [5] that our Theorem 3.1 is essentially the best; i.e., it can not be asymptotically improved.

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